

Supplementary Material for

Long-term air pollution exposure impact on COVID-19 morbidity in China

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Introduction of Poisson regression and negative binomial regression

Poisson regression method is often applied to model count data occurring in different geographic units, in a Poisson model, the probability that an area will experience y events is estimated by

$$P(Y = y) = \exp(-\lambda) \frac{\lambda^y}{y!} \quad y = 0, 1, 2, \dots \quad \lambda > 0 \quad (1)$$

Where, Y represents the event count at geographic unit i during a certain period of time, y denotes a realization of Y . λ is an event occurrence intensity parameter calculated using the total number of events divided by the total number of geographic units. In a Poisson process, it is assumed that the expected value and the variance of the response variable equal its mean, thus

$$\text{Var}(Y) = E(Y) = \lambda \quad (2)$$

Nonetheless, since events are usually unevenly distributed across a region, some studies suggest the intensity parameter λ should be adjusted for this spatial heterogeneity, thereby allowing each geographical unit to have its own intensity parameter λ_i (Cameron and Trivedi, 1986; Beck and Tolnay, 1995;). Under the conditional Poisson assumption, the probability that the i th area will experience y number of events is

$$P(Y_i = y | X) = \exp(-\lambda_i) \frac{\lambda_i^y}{y!} \quad y = 0, 1, 2, \dots \quad \lambda_i > 0 \quad (3)$$

And it is assumed λ_i is a function of multiple explanatory variables, X_{ij} , that describe

each geographical unit

$$E(Y_i | X) = \lambda_i = \exp(\sum_{j=1}^k \beta_j X_{ij}) \quad (4)$$

Where β_j is the estimated effect parameter for explanatory variable X_j .

Because in the presence of over-dispersed count variables, the standard errors of estimated coefficients predicted by a Poisson regression are often underestimated (Cameron and Trivedi, 1986; McCullagh and Nelder, 1989). A regression method that relaxes the mean=variance constraint is thus needed. Based on eq.(4), a negative binomial model can be expressed as

$$E(Y_i | X) = \tilde{\mu}_i = \exp(\sum_{j=1}^k \beta_j X_{ij} + \varepsilon_i) = \mu_i G \quad (5)$$

Where $\tilde{\mu}_i$ is the true count for area i , μ_i is a function of various explanatory variables, and the effect coefficient β_j is estimated using a maximum likelihood approach. and $G = \exp(\varepsilon_i)$, which is a random variable that follows a Gamma distribution. The probability density function of G is

$$f_G(\kappa) = \frac{\kappa^{(1-\alpha)/\alpha} e^{-\kappa/\alpha}}{\alpha^{1/\alpha} \Gamma(1/\alpha)} \quad \kappa > 0 \quad (6)$$

Where $1/\alpha$ and κ represent the shape parameter and the scale parameter of the Gamma distribution, respectively.

The distribution of Y_i around $\tilde{\mu}_i$ is negative binomial, therefore, in a negative binomial model, the probability that geographical unit i will experience y_i events is given by

$$\begin{aligned} P(Y_i = y_i) &= \int_0^\infty P(Y_i = y_i | \tilde{\mu}_i) f(\kappa) d\kappa = \int_0^\infty \frac{(\mu_i \kappa)^{y_i} e^{-\mu_i \kappa} \kappa^{(1-\alpha)/\alpha} e^{-\kappa/\alpha}}{y_i! \alpha^{1/\alpha} \Gamma(1/\alpha)} d\kappa \\ &= \mu_i^{y_i} \frac{\Gamma(y_i + (1/\alpha))}{y_i! \Gamma(1/\alpha)} \frac{(1/\alpha)^{1/\alpha}}{(\mu_i + 1/\alpha)^{y_i + (1/\alpha)}} \quad y = 0, 1, 2, 3 \dots, \alpha, \kappa > 0 \end{aligned} \quad (7)$$

The negative binomial distribution has a mean of μ_i and a variance of

$$Var[Y_i] = \mu_i + \alpha \mu_i^2 \quad (8)$$

As can be seen from eq.(8), NB distribution approaches Poisson distribution as α gradually approaches 0.

The marginal effect of the j th covariate X_j on the expected outcome is given by

$$\delta_j = \left(\frac{1}{\tilde{\mu}_i} \right) \frac{\partial \tilde{\mu}_i}{\partial X_j} = \beta_j \quad (9)$$

It means if we increase one unit of X_j and holding other explanatory variables as constants, the relative rate of change of $\tilde{\mu}_i$ will be β_j .

And the incidence rate ratio (IRR) in a NB model can be derived from $\exp(\beta_j)$, which suggests with one unit increase of X_j from the average of X_j , we expect to see $\exp(\beta_j)$ times the incident events as those with the average of X_j . The reason why $\exp(\beta_j)$ can be interpreted as the incidence rate ratio is explained by UCLA Statistical Consulting Group.

Reference

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Table S1. The single-pollutant model, sensitivity analysis and Poisson model results for COVID-19 incidence rate ratio and CI(95%) associated with per $\mu\text{g}/\text{m}^3$ increase of $\text{PM}_{2.5}$, PM_{10} , NO_2 , CO , SO_2 , and O_3

Regression Model	Pollutant	Regression coefficient	p value	Incidence Rate Ratio(IRR)	95% CI
Single-pollutant model	$\text{PM}_{2.5}$	0.02	<0.001	1.02	(1.01-1.03)
	PM_{10}	0.01	<0.05	1.01	(1.00-1.01)
	SO_2	0.01	0.32	1.01	(0.99-1.02)
	NO_2	0.05	<0.001	1.05	(1.03-1.06)
	CO	0.1	0.68	1.11	(0.66-1.88)
	O_3	-0.02	<0.05	0.98	(0.96-0.99)
Exclude cities in Hubei	$\text{PM}_{2.5}$	0.01	<0.05	1.01	(1.00-1.02)
	PM_{10}	0	<0.05	1.00	(1.00-1.01)
	SO_2	0.02	<0.05	1.02	(1.00-1.03)
	NO_2	0.04	<0.001	1.04	(1.03-1.06)
	CO	0.21	0.35	1.24	(0.79-1.97)
	O_3	-0.01	0.06	0.99	(0.97-1.00)
Exclude number of doctors	$\text{PM}_{2.5}$	0.02	<0.001	1.02	(1.01-1.03)
	PM_{10}	0	0.08	1.00	(1.00-1.01)
	SO_2	0.01	0.42	1.01	(0.99-1.02)
	NO_2	0.04	<0.001	1.05	(1.03-1.06)
	CO	0.07	0.78	1.07	(0.65-1.82)
	O_3	-0.02	<0.05	0.98	(0.96-0.99)
Exclude per capita GDP and average wage	$\text{PM}_{2.5}$	0.02	<0.001	1.02	(1.01-1.03)
	PM_{10}	0	<0.05	1	(1.00-1.01)
	SO_2	0.01	0.45	1.01	(0.99-1.02)
	NO_2	0.04	<0.001	1.04	(1.03-1.06)
	CO	0.04	0.85	1.05	(0.65-1.70)
	O_3	-0.02	<0.05	0.98	(0.96-0.99)
Exclude all weather variables	$\text{PM}_{2.5}$	0.01	0.08	1.01	(1.00-1.02)
	PM_{10}	-0.01	<0.05	1	(0.99-1.00)
	SO_2	-0.01	0.29	0.99	(0.98-1.01)
	NO_2	0.03	<0.001	1.03	(1.01-1.05)
	CO	-0.32	0.20	0.73	(0.43-1.25)
	O_3	-0.03	<0.001	0.97	(0.95-0.98)
Exclude percentage of people over 65-year old	$\text{PM}_{2.5}$	0.02	<0.001	1.02	(1.01-1.03)
	PM_{10}	0	0.07	1.00	(1.00-1.01)
	SO_2	0.01	0.28	1.01	(0.99-1.03)
	NO_2	0.04	<0.001	1.04	(1.03-1.06)
	CO	0.07	0.78	1.07	(0.64-1.82)
	O_3	-0.02	<0.05	0.98	(0.96-0.99)
Exclude mobility	$\text{PM}_{2.5}$	0.02	<0.05	1.02	(1.01-1.04)
	PM_{10}	0	0.45	1.00	(1.00-1.01)

variables	SO ₂	-0.03	<0.001	0.97	(0.95-0.98)
	NO ₂	0.04	<0.001	1.04	(1.02-1.06)
	CO	0.10	0.79	1.11	(0.52-2.36)
	O ₃	-0.01	0.43	0.99	(0.97-1.01)
Poisson Model	PM _{2.5}	0.04	<0.001	1.04	(1.04-1.04)
	PM ₁₀	0.01	<0.001	1.01	(1.01-1.02)
	SO ₂	-0.02	<0.001	0.98	(0.98-0.98)
	NO ₂	0.05	<0.001	1.05	(1.05-1.05)
	CO	0.16	<0.001	1.18	(1.10-1.26)
	O ₃	-0.02	<0.001	0.98	(0.98-0.98)

Table S2. The two-pollutant model results for COVID-19 incidence rate ratio and CI(95%) associated with per $\mu\text{g}/\text{m}^3$ increase of PM_{2.5}, PM₁₀, NO₂, CO, SO₂, and O₃

	Pollutant	coefficient	p value	IRR	95% CI
PM _{2.5}	PM ₁₀	0.04	<0.001	1.05	(1.02-1.07)
	SO ₂	0.02	<0.001	1.02	(1.01-1.03)
	NO ₂	-0.01	0.26	0.99	(0.98-1.01)
	CO	0.02	<0.001	1.03	(1.01-1.04)
	O ₃	0.02	<0.001	1.02	(1.01-1.03)
PM ₁₀	PM _{2.5}	-0.02	<0.05	0.99	(0.97-1.00)
	SO ₂	0.01	<0.05	1.01	(1.00-1.01)
	NO ₂	-0.01	<0.05	0.99	(0.98-1.00)
	CO	0.01	<0.05	1.01	(1.00-1.01)
	O ₃	0.01	<0.05	1.01	(1.00-1.01)
SO ₂	PM _{2.5}	0	0.83	1.00	(0.98-1.02)
	PM ₁₀	0	0.70	1.00	(0.99-1.02)
	NO ₂	-0.01	0.31	0.99	(0.98-1.01)
	CO	0.01	0.38	1.01	(0.99-1.03)
	O ₃	0.01	0.38	1.01	(0.99-1.02)
NO ₂	PM _{2.5}	0.05	<0.001	1.05	(1.03-1.08)
	PM ₁₀	0.06	<0.001	1.06	(1.04-1.08)
	SO ₂	0.05	<0.001	1.05	(1.03-1.07)
	CO	0.06	<0.001	1.06	(1.04-1.08)
	O ₃	0.04	<0.001	1.04	(1.03-1.06)
O ₃	PM _{2.5}	-0.02	<0.05	0.98	(0.97-1.00)
	PM ₁₀	-0.02	<0.05	0.98	(0.97-1.00)
	SO ₂	-0.02	<0.05	0.98	(0.97-1.00)
	CO	-0.02	<0.05	0.98	(0.96-0.99)
	NO ₂	-0.01	0.46	0.99	(0.98-1.01)
CO	PM _{2.5}	-0.48	0.10	0.62	(0.35-1.11)
	PM ₁₀	-0.15	0.58	0.86	(0.48-1.55)
	SO ₂	0	1.00	1.00	(0.57-1.79)
	O ₃	-0.07	0.79	0.93	(0.55-1.59)
	NO ₂	-0.73	<0.05	0.48	(0.28-0.82)

Table S3. A comparison of negative binomial regression and Poisson regression results.

	Regression coefficient (NB model)	Regression coefficient (Poisson model)
<i>Intercept</i>	2.54 ^{***}	1.56 ^{***}
<i>pol</i>	0.02 ^{***}	0.04 ^{***}
<i>den</i>	0.15 [*]	0.04 ^{***}
<i>col</i>	0.02	0.43 ^{***}
<i>old</i>	-0.16 [*]	-0.16 ^{***}
<i>mal</i>	-0.07	-0.05 ^{***}
<i>Lgdp</i>	-0.10	0.26 ^{***}
<i>Lwag</i>	0.14	-0.18 ^{***}
<i>doc</i>	0.30 ^{***}	-0.05 ^{***}
<i>smo</i>	0.06	0.27 ^{***}
<i>sat</i>	-0.39	-1.68 ^{***}
<i>wat</i>	-0.05	0.63 ^{***}
<i>sdt</i>	1.91 ^{***}	3.37 ^{***}
<i>wdt</i>	-0.83 [*]	-1.28 ^{***}
<i>sws</i>	0.03	0.62 ^{***}
<i>wws</i>	0.07	-0.77 ^{***}
<i>mov</i>	1.08 ^{***}	0.35 ^{***}
<i>icm</i>	-0.01	-0.08 ^{***}
<i>AIC</i>	2906	22294

*:p-value <0.05, **: p-value<0.01, ***: p-value<0.001

Figure S1. The annual mean concentrations of PM_{2.5}, PM₁₀, SO₂, NO₂, and O₃ between 2015 and 2019 in mainland China

